# Macdonald polynomials and the multispecies zero range process 

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## overview

(1) Motivation: Macdonald polynomials and interacting particle systems
(2) A new combinatorial formula for $\widetilde{H}_{\lambda}(X ; q, t)$
(3) Multispecies Totally Asymmetric Zero Range Process (mTAZRP)
(4) Markov chain on tableaux
(5) Observables

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- we are interested in studying integrable systems whose exact solutions (stationary distributions) can be expressed in terms of combinatorial formulas or special functions (e.g. Macdonald polynomials)


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Let $\langle$,$\rangle be the standard inner product on \Lambda$. Then $\left\{s_{\lambda}\right\}$ is the unique basis of $\Lambda$ that is:
i. orthogonal with respect to $\langle$,
ii. upper triangular with respect to $\left\{m_{\lambda}\right\}$ :

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s_{\lambda}=m_{\lambda}+\sum_{\mu<\lambda} c_{\mu \lambda} m_{\mu}
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- $s_{\lambda}=\sum_{\sigma} x^{\sigma}$ where $\sigma$ is a semi-standard filling of the Young diagram of shape $\lambda$ E.g. the following are the fillings of shape $(2,1)$ on 3 letters:

$s_{(2,1)}=x_{1}^{2} x_{2}+x_{1}^{2} x_{3}+x_{1} x_{2}^{2}+x_{1} x_{2} x_{3}+x_{1} x_{2} x_{3}+x_{1} x_{3}^{2}+x_{2}^{2} x_{3}+x_{2} x_{3}^{2}=m_{(2,1)}+m_{(1,1,1)}$


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Let $\langle,\rangle_{q, t}$ be the inner product on $\Lambda(q, t)$ given by:

$$
\left\langle p_{\lambda}, p_{\mu}\right\rangle_{q, t}=\delta_{\lambda, \mu} z_{\lambda} \prod_{i \geq 1} \frac{1-q^{\lambda_{i}}}{1-t^{\lambda_{i}}}
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- Example:

$$
P_{(2,1)}(X ; q, t)=m_{(2,1)}+\frac{(1-t)(2+q+t+2 q t)}{1-q t^{2}} m_{(1,1,1)}
$$

## modified Macdonald polynomials $\widetilde{H}_{\lambda}(X ; q, t)$

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- obtained from a normalized form of $P_{\lambda}(X ; q, t)$ by plethystic substitution:

$$
\widetilde{H}_{\lambda}(X ; q, t)=t^{n(\lambda)} J_{\lambda}\left[\frac{X}{1-t^{-1}} ; q, t^{-1}\right]
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where $J_{\lambda}$ is a scalar multiple of $P_{\lambda}$.
Example: $\widetilde{H}_{(2,1)}(X ; q, t)=m_{(3)}+(1+q+t) m_{(2,1)}+(1+2 q+2 t+q t) m_{(1,1,1)}$

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- Haglund-Haiman-Loehr '04 gave formulas for $P_{\lambda}$ and $\widetilde{H}_{\lambda}$ as sums over tableaux with statistics maj and (co)inv:
- $P_{\lambda}(X ; q, t)=\sum_{\substack{\sigma \in \operatorname{dg}(\lambda) \\ \sigma \text { non-attacking }}} q^{\operatorname{maj}(\sigma)} t^{\operatorname{coinv}(\sigma)} x^{\sigma} \prod_{u} \frac{1-t}{1-q^{\operatorname{leg}(u)+1} t^{\operatorname{arm}(u)+1}}$
- $\widetilde{H}_{\lambda}(X ; q, t)=\sum_{\sigma \in \operatorname{dg}(\lambda)} q^{\operatorname{maj}(\sigma)} t^{\operatorname{inv}(\sigma)} x^{\sigma}$


## Combinatorial formulas

- Corteel-M-Williams '18 gave a new formula for $P_{\lambda}$ in terms of multiline queues, which also give formulas for the stationary distribution of the ASEP


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## Combinatorial formulas

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- Garbali-Wheeler '20 gave a formula for $\widetilde{H}_{\lambda}$ using integrability, in terms of colored paths
- Corteel-Haglund-M-Mason-Williams '20 gave a "compressed" formula for $\widetilde{H}_{\lambda}$. Using multiline queues and the plethystic relationship between $\widetilde{H}_{\lambda}$ and $P_{\lambda}$, also conjectured a new formula for $\widetilde{H}_{\lambda}$ with statistics maj and a new statistic quinv:

$$
\widetilde{H}_{\lambda}(X ; q, t)=\sum_{\sigma \in \operatorname{dg}(\lambda)} q^{\mathrm{maj}(\sigma)} t^{\mathrm{quinv}(\sigma)} x^{\sigma}
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## multiline queues and the ASEP

- a multiline queue (MLQ) of type $\lambda, n$ is an arrangement and pairing of balls on a $n \times \lambda_{1}$ lattice, with $\lambda_{j}^{\prime}$ balls in row $j$.



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Angel '08, Ferrari-Martin '07 ( $t=0$ case), Martin '18 (for $q=x_{1}=\cdots=x_{n}=1$ ),
Corteel-M-Williams '18 (general)

- The weight $\operatorname{wt}(M)$ of a multiline queue depends on the parameters $t, q, x_{1}, \ldots, x_{n}$ :
- a string of length $\ell$ corresponds to an ASEP particle of species $\ell$. The labels of the balls in that string are $\ell$


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- skipped balls in the MLQ "correspond" to a coinv statistic in $t$
- wrapping balls in the MLQ correspond to a maj statistic in $q$
- Can be represented by a non-attacking tableau, where each string is mapped to a column of the same height, recording the position of each ball in the MLQ.


## From ASEP to Macdonald polynomials

Theorem (Martin '18, Corteel-M-Williams '18)
The (unnormalized) stationary probability of state $\alpha$ of the mASEP is

$$
\widetilde{\operatorname{Pr}}(\alpha)(t)=\sum_{M: \text { row } 1=\alpha} w t(M)(1, \ldots, 1 ; 1, t)
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## Theorem (Cantini-de Gier-Wheeler '15)

The partition function of $\operatorname{ASEP}(\lambda, n)$ is a specialization of the Macdonald polynomial:

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P_{\lambda}(1, \ldots, 1 ; 1, t)=\mathcal{Z}_{\lambda, n}(t)=\sum_{\alpha \in S_{n} \cdot \lambda} \widetilde{\operatorname{Pr}}(\alpha)(t)
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P_{\lambda}\left(x_{1}, \ldots, x_{n} ; q, t\right)=\sum_{M \in \operatorname{MLQ}(\lambda, n)} w t(M)\left(x_{1}, \ldots, x_{n} ; q, t\right)
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$$

This formula essentially coincides with that of Lenart '09 for $\lambda$ with distinct parts.

## Example for $P_{(2,1)}\left(x_{1}, x_{2}, x_{3} ; q, t\right)$

$$
P_{(2,1)}\left(x_{1}, x_{2}, x_{3} ; q, t\right)=m_{(2,1)}+\frac{(2+t+q+2 q t)(1-t)}{\left(1-q t^{2}\right)} m_{(1,1,1)}
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## From multiline queues to a new formula for $\widetilde{H}_{\lambda}$

- Recall: $\widetilde{H}_{\lambda}(X ; q, t)$ is obtained from the integral form of $P_{\lambda}$ via plethysm:

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\begin{aligned}
\widetilde{H}_{\lambda}(X ; \boldsymbol{q}, t) & =t^{n(\lambda)} J_{\lambda}\left[\frac{X}{1-t^{-1}} ; \boldsymbol{q}, t^{-1}\right] \\
& =f_{\lambda}(\boldsymbol{q}, t) P_{\lambda}\left(x_{1}, x_{1} t^{-1}, x_{1} t^{-2}, \ldots, x_{2}, x_{2} t^{-1}, x_{2} t^{-2}, \ldots ; \boldsymbol{q}, t^{-1}\right)
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- this leads to a new "queue inversion" statistic for $t$ that we call quinv (Corteel-Haglund-M-Mason-Williams '20, Ayyer-M-Martin '21)
- the resulting objects are of the same flavor as multiline queues, except that multiple balls are allowed at each location. (This translates to removing the "non-attacking" condition from the corresponding tableaux)


## tableaux formulas: notation and statistics

- $\operatorname{dg}(\lambda)$ (the diagram of $\left.\lambda=\left(\lambda_{1}, \ldots, \lambda_{k}\right)\right)$ consists of $k$ bottom justified columns with $\lambda_{i}$ boxes, from left to right



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$$
\lambda=(4,3,3,1)
$$

$\sigma=$| 4 |  |  |  |
| :--- | :--- | :--- | :---: |
| 2 | 2 | 4 |  | |  |  |  |
| :--- | :--- | :--- |
| 3 | 1 | 1 |

- a tableau of type $(\lambda, n)$ is a filling $\sigma: \operatorname{dg}(\lambda) \rightarrow[n]$ of the cells


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| $\sigma=$ |  |  |  |
| :---: | :---: | :---: | :---: |
|  | 2 | 4 |  |
| 3 | 1 | 1 |  |
| 2 | 3 | 3 | 4 |

- a tableau of type $(\lambda, n)$ is a filling $\sigma: \operatorname{dg}(\lambda) \rightarrow[n]$ of the cells
- $\operatorname{inv}(\sigma)$ is the number of inversions in the configuration

$$
\begin{array}{|l|l|l|}
\hline x & \cdots & z \\
y & & \text { where } x<y<z(\text { cyclically } \bmod n) \\
\hline y &
\end{array}
$$

## tableaux formulas: notation and statistics

- $\operatorname{dg}(\lambda)$ (the diagram of $\left.\lambda=\left(\lambda_{1}, \ldots, \lambda_{k}\right)\right)$ consists of $k$ bottom justified columns with $\lambda_{i}$ boxes, from left to right

$$
\sigma=
$$

- a tableau of type $(\lambda, n)$ is a filling $\sigma: \operatorname{dg}(\lambda) \rightarrow[n]$ of the cells
- $\operatorname{inv}(\sigma)$ is the number of inversions in the configuration

$$
\begin{array}{|l|l|}
\hline x & \cdots \\
y & z
\end{array} \quad \text { where } x<y<z(\text { cyclically } \bmod n)
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## Theorem (Haglund-Haiman-Loehr '05)

The modified Macdonald polynomial is given by

$$
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$$

## a new statistic: queue-inversion

$$
\left.\sigma=\right) 4 .
$$

- an $L$-triple is a triple of cells in the configuration:

$$
\begin{array}{|c|c|cc|c|}
\hline x \\
\hline y & \cdots & \begin{array}{c} 
\\
\hline
\end{array} \quad \text { or } \quad \begin{array}{|c}
y \\
\hline
\end{array} & \cdots & z \\
\hline
\end{array}
$$

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& 4 \\
& \hline
\end{aligned}
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\hline
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## a new statistic: queue-inversion

$\sigma=$| 4 |  |  |  |
| :--- | :--- | :--- | :--- |
| 2 | 2 | 4 |  |
| 3 | 1 | 1 |  |
| 2 | 3 | 3 | 4 |

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Let $\lambda$ be a partition. The modified Macdonald polynomial equals

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| :--- | :--- | :--- |
| 2 | 2 | 4 | \(\left.\begin{array}{|l|l|}\hline \& 1 <br>

2 \& 1\end{array}\right)\)

$$
\text { quinv }(\sigma)=4
$$

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## Example: $\tilde{H}_{(2,1)}(X ; q, t)$

$\widetilde{H}_{(2,1)}\left(x_{1}, x_{2} ; q, t\right)=m_{(3)}+(1+t+q) m_{(2,1)}+(1+2 t+2 q+q t) m_{(1,1,1)}$

- (AMM) $\quad \widetilde{H}_{\lambda}(X ; q, t)=\sum_{\sigma: \mathrm{dg}(\lambda) \rightarrow \mathbb{Z}_{+}} q^{\operatorname{maj}(\sigma)} t^{q \mathrm{qunv}(\sigma)} x^{\sigma}$

- (HHL) $\quad \widetilde{H}_{\lambda}(X ; q, t)=\sum_{\sigma: \operatorname{dg}(\lambda) \rightarrow \mathbb{Z}_{+}} q^{\operatorname{mi}(\sigma)} t^{\min (\sigma)} x^{\sigma}$

| 1 | 2 | 1 | 1 | 1 | 2 | 3 | 2 | 1 | 3 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 11 | 11 | 12 | 21 | 23 | 13 | 12 | 31 | 32 | 21 |
| $m_{3}$ | $q m_{21}$ | $m_{21}$ | $t m_{21}$ | $m_{111}$ | $q m_{111}$ | $q m_{111}$ | $t m_{111}$ | $t m_{111}$ | qt $m_{111}$ |

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$$
\left.\begin{array}{cccccccccccccccccc}
1 & & 2 & & 1 & 1 & 1 & 2 & 3 & 2 & 1 & 3 & \\
1 & 1 & 1 & 1 & 1 & 2 & 2 & 1 & 2 & 3 & 1 & 3 & 1 & 2 & 3 & 1 & 3 & 2
\end{array}\right] 2 \begin{gathered}
1 \\
m_{3}
\end{gathered}
$$

- while the inv and quinv statistics appear very similar, there does not seem to be an easy way to go from one to the other - is there a bijective proof?


## Motivation



What is the analogous interacting particle system whose partition function is a specialization of $\widetilde{H}_{\lambda}$ ?

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## totally asymmetric zero range processes (TAZRP)

- continuous-time stochastic processes (Spitzer '70), can be defined on arbitrary graphs. In our case, we have a circular lattice with $n$ sites.



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$$
\begin{gathered}
\text { Here, } n=5, k=7 \\
\tau=(11|\cdot| 111|1| 1)
\end{gathered}
$$

- simplest case: there are $k$ indistinguishable particles, moving counter-clockwise. A configuration $\tau=\left(\tau_{1}, \ldots, \tau_{n}\right)$ is any allocation of the $k$ particles on the $n$ sites.


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$$
\leftarrow 3,3,1
$$

$$
\begin{gathered}
\text { Here, } n=5, k=7 \\
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- multispecies variant: we now allow different particle types, labeled by integers (particles of the same type are still indistinguishable)
- Kuniba-Maruyama-Okado (2015+) (and others) have studied many multispecies variants of the TAZRP. All of these are integrable! The version we will describe was first studied by Takayama '15


## the mTAZRP: states

- Fix a (circular 1D) lattice on $n$ sites and a partition $\lambda=\left(\lambda_{1} \geq \cdots \geq \lambda_{k}>0\right)$ for the particle types



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- Fix a (circular 1D) lattice on $n$ sites and a partition $\lambda=\left(\lambda_{1} \geq \cdots \geq \lambda_{k}>0\right)$ for the particle types
- $\operatorname{TAZRP}(\lambda, n)$ is a Markov chain whose states are multiset compositions $\tau$ of type $\lambda$, with $n$ (possibly empty) parts



## the mTAZRP: transition rates

- Each particle is equipped with an exponential clock. Transitions are jumps from site $j$ to site $j+1$


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- Each particle is equipped with an exponential clock. Transitions are jumps from site $j$ to site $j+1$
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- For $1 \leq j \leq n$ and $k \in \lambda$, call $f_{j}(k)$ the rate of the jump of particle $k$ from site $j$ to site $j+1$. If site $j$ has $d$ particles larger than $k$ and $c$ particles of type $k$, then

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f_{j}(k)=x_{j}^{-1} t^{d} \sum_{u=0}^{c-1} t^{u}
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$$

For example: If site $j$ contains the particles $\{4,3,3,1,1,1\}$, then:

$$
\begin{array}{lll}
k=1: & d=3, & c=3, \\
k=3: & d=1, & c=2, \\
k=4: & d=0, & c=1,
\end{array}
$$

## Lumping of tableaux to mTAZRP

- Given a filling $\sigma$, read the state $\tau \in \operatorname{TAZRP}(\lambda, n)$ from the bottom row of $\sigma$ as follows:
$\tau_{j}$ is the multiset $\left\{\lambda_{i}: \sigma(1, i)=j\right\}$


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$\tau_{j}$ is the multiset $\left\{\lambda_{i}: \sigma(1, i)=j\right\}$
- For example, for $\lambda=(2,1,1)$ and $n=3$, the following are all the tableaux that correspond to the state $\tau=(21|\cdot| 1)$ :

| 1 |  |  | 2 |  | 3 |  |  | 1 |  |  | 2 |  |  | 3 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 1 |  | 1 | 13 | 1 | 1 | 3 | 1 | 3 | 1 | 1 | 3 | 1 | 1 | 3 | 1 |

## TAZRP probabilities and tableaux

## Theorem (Ayyer-M-Martin '21)

Fix $\lambda, n$. The (unnormalized) stationary probability of $\tau \in \operatorname{TAZRP}(\lambda, n)$ is

$$
\widetilde{\operatorname{Pr}}(\tau)=\sum_{\substack{\sigma: \mathrm{dg}(\lambda) \rightarrow[n] \\ \sigma \text { has type } \tau}} x^{\sigma} t^{\mathrm{quinv}(\sigma)} .
$$

## Corollary

The so-called partition function of $\operatorname{TAZRP}(\lambda, n)$ is

$$
\mathcal{Z}_{\lambda, n}\left(x_{1}, \ldots, x_{n} ; t\right)=\widetilde{H}_{\lambda}\left(x_{1}, \ldots, x_{n} ; 1, t\right)
$$

## an example for $\lambda=(2,1,1)$ and $n=2$

The stationary distribution is:

Example computation for $(21 \mid 1)$ :

| 1 |  |  |
| :--- | :--- | :--- |
| 1 | 1 | 2 |\(: \begin{aligned} \& <br>

\& t^{2},\end{aligned}\)

$$
\begin{array}{|l|l|l}
\hline 2 & & \\
\hline 1 & 1 & 2 \\
\hline
\end{array}
$$

$$
\begin{array}{|l|l|l}
\hline 1 & & \\
\hline 1 & 2 & 1 \\
\hline
\end{array}
$$

$$
\begin{array}{|l|l|l}
\hline 2 & & \\
\hline 1 & 2 & 1 \\
\hline
\end{array}
$$

the total is:

$$
\widetilde{\operatorname{Pr}}(21 \mid 1)=x_{1}^{2} x_{2}\left(t x_{1}+x_{2}\right)(1+t) .
$$

## why queue inversions? multiline diagrams

The tableaux are actually representing a queueing system which is an arrangement of lattice paths/strings: the lattice paths are representing the coupling of individual single species TAZRPs

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## a Markov chain on tableaux: notation

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- each cell $c$ such that $\sigma(\operatorname{South}(c)) \neq \sigma(c)$ is equipped with an exponential clock with rate

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- If $c=(1, j)$ is in the bottom row, then $\operatorname{arm}(\sigma, c)$ is equal to the number of particles larger than or equal to $\lambda_{j}$ at site $\sigma(c)$ of the corresponding state of the TAZRP. Thus $f(\sigma, c)$ is equal to the rate of the correponding TAZRP jump.


## a Markov chain on tableaux: transitions

A transition $M_{c}$ triggered by a cell $c$ : if $\sigma(c) \neq \sigma($ South $(c))$, take the maximal contiguous (cyclically) increasing chain of cells weakly above $c$ in its column, and increment the content of each cell by 1 . (This is sometimes called a ringing path)


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| 4 |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| 2 | 1 |  |  | $\xrightarrow{t x_{3}^{-1}}$ |
| 1 | 1 |  |  |  |
| 3 | 4 | 3 |  |  |
| 3 | 3 | 4 | 1 |  |



## Theorem (Ayyer-M-Martin '21)

The stationary distribution of the Markov process on the tableaux is

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- if $c=(1, j)$ is in the bottom row, the rate $f(\sigma, c)$ matches the transition rate $f_{\sigma(c)}\left(\lambda_{j}\right)$ of the corresponding particle in the TAZRP.
- (when $\lambda$ has repeated parts, we need to do some more work!)


## a Markov chain on tableaux: proof

$$
M(\sigma)=\left\{M_{c}(\sigma): c \in \operatorname{dg}(\lambda), \sigma(c) \neq \sigma(\operatorname{South}(c))\right\}
$$

| 2 |  |
| :--- | :--- |
| 1 | 2 |

## a Markov chain on tableaux: proof

$$
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$$



## a Markov chain on tableaux: proof

$$
\begin{aligned}
M(\sigma) & =\left\{M_{c}(\sigma): c \in \operatorname{dg}(\lambda), \sigma(c) \neq \sigma(\operatorname{South}(c))\right\} \\
R(\sigma) & =\left\{\sigma^{\prime} \in \operatorname{PQT}(\lambda, n): T \in M\left(\sigma^{\prime}\right)\right\}
\end{aligned}
$$



## balance equation

if each $\sigma \in \operatorname{dg}(\lambda) \rightarrow[n]$ satisfies:

$$
\sum_{\sigma^{\prime} \in R(\sigma)} w t\left(\sigma^{\prime}\right) \operatorname{rate}\left(\sigma^{\prime} \rightarrow \sigma\right)=\mathrm{wt}(\sigma) \sum_{\sigma^{\prime} \in M(\sigma)} \operatorname{rate}\left(\sigma \rightarrow \sigma^{\prime}\right),
$$

then the stationary distribution of the M.C. on tableaux is $\mathrm{wt}(\sigma)$.

## Observables: partition function

- The partition function of the mTAZRP of type $\lambda, n$ is defined to be the normalizing constant, or the sum of the unnormalized stationary probabilities:

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Z_{\lambda, n}=\sum_{\tau} \tilde{\operatorname{Pr}}(\tau)
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- We have

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\begin{aligned}
Z_{\lambda, n} & =\widetilde{H}_{\lambda}\left(x_{1}, \ldots, x_{n} ; 1, t\right)=\prod_{j=1}^{\lambda_{1}} \widetilde{H}_{\left(1^{\lambda_{j}^{\prime}}\right)}\left(x_{1}, \ldots, x_{n} ; 1, t\right) \\
& =\prod_{j=1}^{\lambda_{1}} \sum_{\mu \vdash \lambda_{j}^{\prime}}\left[\begin{array}{c}
\lambda_{j}^{\prime} \\
\mu
\end{array}\right]_{t} m_{\mu}\left(x_{1}, \ldots, x_{n}\right) .
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Notice that $Z$ might have extra factors, e.g. when the gcd of the probabilities $\widetilde{\operatorname{Pr}}(\tau)$ is not 1 .

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$$

Notice that $Z$ might have extra factors, e.g. when the gcd of the probabilities $\widetilde{\operatorname{Pr}}(\tau)$ is not 1 .

- Sanity check: from the point of view of the TAZRP, having three species of particles labeled $1,2,3$ is the same process as having three species labeled $2,13,27$. Thus we should expect their stationary probabilities to be proportional.


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- The partition function of the mTAZRP of type $\lambda, n$ is defined to be the normalizing constant, or the sum of the unnormalized stationary probabilities:

$$
Z_{\lambda, n}=\sum_{\tau} \tilde{\operatorname{Pr}}(\tau)
$$

- We have

$$
\begin{aligned}
Z_{\lambda, n} & =\widetilde{H}_{\lambda}\left(x_{1}, \ldots, x_{n} ; 1, t\right)=\prod_{j=1}^{\lambda_{1}} \widetilde{H}_{\left(1^{\lambda_{j}^{\prime}}\right)}\left(x_{1}, \ldots, x_{n} ; 1, t\right) \\
& =\prod_{j=1}^{\lambda_{1}} \sum_{\mu \vdash \lambda_{j}^{\prime}}\left[\begin{array}{c}
\lambda_{j}^{\prime} \\
\mu
\end{array}\right]_{t} m_{\mu}\left(x_{1}, \ldots, x_{n}\right)
\end{aligned}
$$

Notice that $Z$ might have extra factors, e.g. when the gcd of the probabilities $\widetilde{\operatorname{Pr}}(\tau)$ is not 1 .

- Sanity check: from the point of view of the TAZRP, having three species of particles labeled $1,2,3$ is the same process as having three species labeled $2,13,27$. Thus we should expect their stationary probabilities to be proportional.
- At the very least, we need $\widetilde{H}_{(2,13,27)}\left(x_{1}, \ldots, x_{n} ; 1, t\right)$ to be divisible by $\widetilde{H}_{(1,2,3)}\left(x_{1}, \ldots, x_{n} ; 1, t\right)$. This is indeed true, since $(3,2,1)^{\prime} \subset(27,13,2)^{\prime}$.


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\begin{gathered}
\text { Here, } n=5, m=7 \\
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## Proposition (Current for the single species TAZRP)

For the single-species TAZRP on $n$ sites with $m$ particles, the current is given by

$$
J=[m]_{t} \frac{\widetilde{H}_{\left\langle 1^{m-1}\right\rangle}\left(x_{1}, \ldots, x_{n} ; 1, t\right)}{\widetilde{H}_{\left\langle 1^{m}\right\rangle}\left(x_{1}, \ldots, x_{n} ; 1, t\right)}
$$

## Current

## Theorem (Ayyer-M-Martin '22+)

Let $\lambda=\left(1^{m_{1}}, \ldots, k^{m_{k}}\right)$, and let $1 \leq j \leq k$. The current of the particle of type $j$ of the TAZRP of type $\lambda$ on $n$ sites is given by

$$
J=\left[m_{j}+\cdots+m_{k}\right]_{t} \frac{\widetilde{H}_{\left(1^{m_{j}+\cdots+m_{k}-1}\right)}}{\widetilde{H}_{\left(1^{m_{j}+\cdots+m_{k}}\right)}}
$$

$$
-\left[m_{j+1}+\cdots+m_{k}\right]_{t} \frac{\widetilde{H}_{\left(1^{m_{j+1}+\cdots+m_{k}-1}\right)}}{\widetilde{H}_{\left(1^{m_{j+1}+\cdots+m_{k}}\right)}}
$$

## Densities

- Take $\operatorname{TAZRP}(\lambda, n)$ with content $\lambda=\left(1^{m_{1}}, 2^{m_{2}}, \ldots, k^{m_{k}}\right)$.
- Define $z_{j}^{(\ell)}$ to be the random variable counting the number of particles of type $\ell$ at site $j$ in a configuration of $\operatorname{TAZRP}(\lambda, n)$.
- Denote the expectation in the stationary distribution by $\left\langle z_{j}^{(\ell)}\right\rangle$.


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## Proposition (Translation invariance)

Suppose $\left\langle z_{1}^{(\ell)}\right\rangle=r\left(x_{1}, \ldots, x_{n}\right)$. Then for any $j$,

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\left\langle z_{j}^{(\ell)}\right\rangle=r\left(x_{j}, \ldots, x_{n}, x_{1}, \ldots, x_{j-1}\right)
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We begin with the special case of $\lambda=1^{m}$.

## Theorem (Densities for the single species TAZRP)

The density at site 1 on $\operatorname{TAZRP}\left(1^{m}, n\right)$ is given by

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- In particular, when $x_{1}=\cdots=x_{n}=1$, the density is $\left\langle z_{1}^{(1)}\right\rangle=\frac{m}{n}$.


## Densities

## Theorem (Ayyer-M-Martin '22+)

For $1 \leq \ell \leq k$, the density of the $\ell$ 'th species at site 1 is given by

$$
\left\langle z_{1}^{(\ell)}\right\rangle=x_{1} \partial_{x_{1}} \log \left(\frac{\widetilde{H}_{\left(1^{m_{\ell}}+\cdots+m_{k}\right)}\left(x_{1}, \ldots, x_{n} ; 1, t\right)}{\widetilde{H}_{\left(1^{m_{\ell+1}+\cdots+m_{k}}\right)}\left(x_{1}, \ldots, x_{n} ; 1, t\right)}\right) .
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## Corollary

$\left\langle z_{1}^{(\ell)}\right\rangle$ is symmetric in the variables $\left\{x_{2}, \ldots, x_{n}\right\}$.

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## Corollary

$\left\langle z_{1}^{(\ell)}\right\rangle$ is symmetric in the variables $\left\{x_{2}, \ldots, x_{n}\right\}$.
Proof via coloring argument:

- true for base case $\lambda=\left(1^{m}\right)$
- transitions of particles of species $\ell, \ldots, k$ at site 1 are independent of the number of lower species particles at site 1 . Thus we can ignore the particles of types $1, \ldots, \ell-1$.
- the density of particles of species $\ell, \ldots, k$ at site 1 is equivalent to the density at site 1 of a TAZRP of type $\lambda=\left(1^{m_{\ell}+\cdots+m_{k}}\right)$.
- to isolate species $\ell$ we subtract the density of species $\ell+1, \ldots, k$ from the density of species $\ell, \ldots, k$.


## Local correlations

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- We consider two kinds of local correlations:
- Let $\mathbb{P}_{\lambda, n}(\bar{w})$ be the stationary probability of having exactly the content $w_{1}, \ldots, w_{\ell}$ on sites $1, \ldots, \ell$.
- Let $\mathbb{P}_{\lambda, n}(\hat{w})$ be the stationary probability of having at least the content $w_{1}, \ldots, w_{\ell}$ on sites $1, \ldots, \ell$.


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- Example: let $\lambda=(2,2,1,1), n=4, \ell=2$, and $w=(2 \mid 1)$.
- Configurations contributing to $\mathbb{P}_{\lambda, n}(\bar{w})$ are

$$
(2|1| 12 \mid \cdot), \quad(2|1| 1 \mid 2), \quad(2|1| 2 \mid 1), \quad(2|1| \cdot \mid 12)
$$

- Additional configurations contributing to $\mathbb{P}_{\lambda, n}(\hat{w})$ are
$(12|1| 2 \mid \cdot)$,
(2|11|2|•)
$(22|1| 1 \mid \cdot)$
(2|12|1|•)
(12|1|•|2),
(2|11|•|2)
(22|1|•|1),
$(2|12| \cdot \mid 1), \quad(122|1| \cdot \mid \cdot)$,
$(2|112| \cdot \mid \cdot), \quad(22|11| \cdot \mid \cdot)$,
(12|12|•|)


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Theorem (Ayyer-M-Martin '22+)
Both $\mathbb{P}_{\lambda, n}(\bar{w})$ and $\mathbb{P}_{\lambda, n}(\hat{w})$ are symmetric in the variables $\left\{x_{\ell+1}, \ldots, x_{n}\right\}$.

## final remarks

- Explicit bijection from the inv to the quinv statistic?

- Can we find a dynamical process that incorporates the $q$ as a parameter?

This seems difficult because

- We lose factorization of $\widetilde{H}_{\lambda}$
- We lose translation invariance
- Using multiline queues (for the ASEP on a circle),

Corteel-Haglund-M-Mason-Williams '20 defined quasisymmetric Macdonald polynomials which refine $P_{\lambda}$. Can we use a parallel construction to define an interesting family of quasisymmetric polynomials that refine $\widetilde{H}_{\lambda}$ ?

- Same as above, but for nonsymmetric Macdonald polynomials

- Modified Macdonald polynomials and the multispecies zero range process: I, (with A. Ayyer and J. B. Martin), arXiv:2011.06117
- A Markov chain on tableaux that projects to the multispecies TAZRP, and applications, (with A. Ayyer and J. B. Martin), in preparation

